

**Social Norms and Rank-Based Nudging: Changing Willingness to Pay for Healthy Food**

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Social Norms and Rank-Based Nudging: Changing Willingness to Pay for Healthy Food

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### Abstract

People's evaluations in the domain of healthy eating are at least partly determined by the choice context. We systematically test reference level and rank-based models of relative comparisons against each other and explore their application to social norms nudging, an intervention that aims at influencing consumers' behavior by addressing their inaccurate beliefs about their consumption relative to the consumption of others. Study 1 finds that the rank of a product or behavior amongst others in the immediate comparison context, rather than its objective attributes, influences its evaluation. Study 2 finds that when a comparator is presented in isolation the same rank-based process occurs based on information retrieved from memory. Study 3 finds that telling people how their consumption ranks within a normative comparison sample increases willingness to pay for a healthy food by over 30% relative to the normal social norms intervention that tells them how they compare to the average. We conclude that social norms interventions should present rank information (e.g., "you are in the most unhealthy 10% of eaters") rather than information relative to the average (e.g., "you consume 500 calories more than the average person").

*Keywords:* food perception; healthy eating; Decision by Sampling; Range Frequency Theory; social norms marketing

### Social Norms and Rank-Based Nudging: Changing Willingness to Pay for Healthy Food

Despite extensive nutrition information campaigns and food labeling policies (e.g., Burton, Creyer, Kees, & Huggins, 2006; Freeland-Graves & Nitzke, 2002), the prevalence of health-related issues arising from poor dietary choices is on the rise. It has been suggested that one factor contributing to food overconsumption is the relative nature of people's evaluation of food products and their healthiness, which can lead to the same foods being appraised differently depending on the other food options available in the decision-making context (Chandon & Wansink, 2012; Chernev, 2011; Geier, Rozin, & Doros, 2006; Sharpe, Staelin, & Huber, 2008; Wansink, Just, & Payne, 2009).

In the present study, we first contribute to this literature using the rank principle as embodied in rank-based models such as Range Frequency Theory (RFT; Parducci, 1965) and Decision by Sampling (DbS; Stewart, Chater, & Brown, 2006), which have previously been applied in cognitive and social psychology. We extend the models here to both the consumption and the broader consumer research (cf. also Niedrich, Sharma, & Wedell, 2001; Niedrich, Weathers, Hill, & Bell, 2009). In doing so we explain how people's evaluations and purchase of products involve integration of information about a product with information on rival products present in the choice context or retrieved from memory. Second, in direct tests across three studies we advance theory in the field of consumers' food choice by finding support for the predictions of rank-based models rather than the extant dominant model (based on reference level; Helson, 1947). Third, we show how to improve social norms interventions, as commonly used in social marketing and public health, through a minor reframing of the interventional messages to target people's natural ways of processing information as suggested by DbS and RFT.

### **Contextual Influence on Food Evaluation and Choice**

How do people assess the healthiness of their own diet or of the food they purchase? If people make judgments purely in *absolute* terms, then the same food product will be evaluated in the same way regardless of other food options available. However, people perform poorly when making judgments based on the absolute magnitude of food or drink portions. For example, if the size of a food portion is doubled, people usually report this increase to be only around 50 to 70%. This bias is associated with higher likelihood of food overconsumption, especially because of the increasingly widespread offers of larger package sizes which are more profitable for food marketers (Chandon & Wansink, 2012).

In contrast to the absolute account, evidence from both cognitive (e.g., Stewart, Brown, & Chater, 2005) and consumer psychology (e.g., Sharpe et al., 2008) has suggested that people evaluate size in *relative* terms – that is, people are highly sensitive to the *context* in which an evaluation (or a choice) is made. An individual product (e.g., a cereal bar, a ready meal) is evaluated with reference to other products (e.g., other cereal bars, other ready meals) available in the decision-making context – these can be labeled as effects of the *immediate* context. In real world scenarios, it has been observed that people's choices about food are influenced by the sizes of available options. For example, in a menu with three drink options (12, 16, and 24 fl. oz.), removing the 12 fl. oz. option leads around 25% of consumers who previously chose a 16 fl. oz. drink to switch a larger one; the 16 fl. oz. option is chosen less when it becomes the smallest available (Sharpe et al., 2008). Similar work has shown that such context effects can lead to up to 30% differences in how much food is consumed (Wansink et al., 2009). Context also affects taste; a sweetness that is judged 'most pleasant' in one context may be deemed unpleasantly sweet in another (Riskey, Parducci, & Beauchamp, 1979). Finally, adding a healthy option can reduce the estimate of the number of calories in an unhealthy food.

Similarly, social norms approaches propose that people are influenced by another type of contextual influence—the *interpersonal* context—which are the norms that people derive when they compare themselves to others (Campbell, 1964; Henrich et al., 2001; Wansink, 2004). A person consumes more or less food depending on how much other people around them are consuming (Herman, Roth, & Polivy, 2003; Wansink et al., 2009). Also, people evaluate their own eating habits by comparing them to what they think other people (especially relevant others, like friends and family) consume (similarly to what people do when they decide whether to pay taxes or not; Posner, 2000).

### **Models of Relative Judgment**

Thus, immediate and interpersonal contexts lead to a variety of effects including; (a) biased estimation of information about a given food (e.g., how many calories it contains, consequences of consumption), and (b) changes in behavior (e.g., product purchase, amount consumed). These context effects reflect cognitive processes of judgment and choice (Wansink et al., 2009).

Reference-level accounts, derived from Adaptation Level Theory (ALT; Helson, 1964) suggest that judgments of a stimulus depend on how its magnitude compares to a single reference point. For example, a person may judge the healthiness of a sandwich based on how its fat, calorie, sugar, and salt content compares to that of a single ‘typical’, ‘average’, ‘prototypical’, or ‘reference’ sandwich. People are hypothesized to form an internalized reference point (the ‘adaptation level’) derived from both the current context (sandwiches on the menu) and prior context retrieved from memory (previously encountered sandwiches). The healthiness of the product under consideration will then be judged against the adaptation level, which here is operationalized as the mean of remembered quantities or items available in the decision-making context. The theory proposes that people continually update their adaptation level, so that new relevant information (e.g., a particularly calorific sandwich on a

menu) will shift the adaptation level upwards or downwards depending on its size. Reference-level effects have been observed in different domains. For instance, recent models suggest that income is evaluated relative to a reference level (Clark & Oswald, 1996), although where such demonstrations have been made they have typically not controlled for the predictions of alternative rank-based models of judgment.

The Decision by Sampling (DbS; Stewart et al., 2006) model suggests that when people make judgments about the magnitude of a target they are influenced by the *rank* principle – that is, people appear to be sensitive to how a quantity ranks within a given context. DbS extends earlier accounts such as Range Frequency Theory (Parducci, 1965) both in specifying the underlying process giving rise to rank-based comparison and in emphasizing the role of samples drawn from long-term memory in forming a context of comparison.

Consider how a person would evaluate the healthiness of a sandwich that has 560 calories. DbS suggests that when a judgment is made, people sample from memory and from the decision-making context in order to make a relative judgment based on binary ordinal comparison. People would retrieve calorie-content information from memory and look at the labels of other sandwiches on display; for example, one might think of half a dozen sandwiches that contained fewer calories, but only one or two with higher calorific content. The subjective evaluation of the healthiness of the sandwich is assumed to be directly determined by its relative rank value within this sample. As a consequence, if different samples are retrieved from memory or available in the decision-making context, the same content might be regarded very differently. DbS provides a plausible description of the psychological processes underlying judgment formation. RFT works efficiently as a descriptive account of judgments in context; however, it seems unlikely that people would keep in memory the entire distribution of values they had encountered in a given domain as would be needed to compute relative rank accurately (although see Parducci, 1992).

However, the present paradigm was not designed to test the predictions of DbS against those of RFT, and thus we will refer to the broader class of rank-based models.

Rank effects have been observed for the evaluation of different entities, ranging from psychophysical stimuli (Parducci & Perrett, 1971) to cognitive and social quantities such as satisfaction with body image (Wedell, Santoyo, & Pettibone, 2005), wages (Brown, Gardner, Oswald, & Qian, 2008; Hagerty, 2000), health and well-being (Boyce, Brown, & Moore, 2010; Boyce & Wood, in press; Wood, Boyce, Moore, & Brown, 2012), gratitude (Wood, Brown, & Maltby, 2011), satisfaction with educational provision (Brown, Wood, Ogden, & Maltby, 2015), fairness of sentencing (Aldrovandi, Wood, & Brown, 2013), indebtedness (Aldrovandi, Wood, Maltby, & Brown, in press), and perception of health risks due to alcohol consumption (Wood, Brown, & Maltby, 2012). However, the implications of the rank-based models have not been explored within the literature on food evaluation or within the marketing and consumer research literature more generally (although see Niedrich et al., 2001; Niedrich et al., 2009). In these literatures, the reference level account still predominates.

### **Social Norms and Practical Implications**

Rank-based and reference-level accounts both address context effects in judgment and decision-making, but they make different predictions about people's evaluations of food products and have different implications for the design of interventions to nudge healthy food consumption.

Recently behavioral 'nudging', a method for guiding and influencing consumer's behavior by shaping the environment but without unduly restricting their freedom of choice (Thaler & Sunstein, 2008), has been implemented within the social norms framework (Agostinelli, Brown, & Miller, 1995). Indeed, many behavioral 'nudges' provide social norm information and have been found to induce behavior change in a variety of contexts such as

energy consumption (Ayres, Raseman, & Shih, 2013; Schultz, Nolan, Cialdini, Goldstein, & Griskevicius, 2007) and recycling (Goldstein, Cialdini, & Griskevicius, 2008). Social norm approaches propose that people evaluate and choose their own behavior at least partly with reference to their beliefs, which are often inaccurate (Prentice & Miller, 1993), about what other people do. To address these inaccurate beliefs, norm-based nudging interventions typically expose people to normative information about what other people *actually* do. This information can cause an individual's behavior to move towards the social norm, because of a preference for social conformity (Festinger, 1957) or because of an assumption that others' behavior is informed by additional knowledge that the individual does not have (Bikhchandani, Hirshleifer, & Welch, 1992).

Beneficial effects are however sometimes small in magnitude (Loewenstein & Ubel, 2010), population-specific (Beshears, Choi, Laibson, Madrian, & Milkman, 2011; Costa & Kahn, 2013) or absent (Russell, Clapp, & DeJong, 2005; Werch et al., 2000). Informed by the rank-based models, we hypothesize that the behavior-changing effect of social norms will depend on precisely what information is presented.

Many interventions provide information about *mean* levels of others' behavior (Moreira, Smith, & Foxcroft, 2009; Schultz et al., 2007); such interventions can be seen as deriving from reference-level accounts (Helson, 1964). Applied to social norms, reference-level approaches could provide support for the idea that people judge the level of their own behavior (e.g., amount of energy consumption) with reference to the mean of the social distribution of levels of such behavior.

However, the contextual effects reviewed in the previous section suggest that the provision of mean-based social norms may not resonate with people's natural ways of making subjective judgments about their own behavior. Thus, motivated by rank-based models, we examine whether rank-based nudging ("90% of people consume less chocolate

than you do”) will be more effective than mean-based nudging (“you eat 5 bars of chocolate per week; on average, other people consume 3 bars per week”) in influencing consumers’ choices about food products. DbS suggests that nudging based on social rank information might be more influential than mean-based nudging because it is beliefs about rank that determine judgments. Intuitively, telling someone only the average level of a behavior provides relatively impoverished social norm information: A consumption of five chocolate bars a week relative to a mean of three might place a consumer in the top 10% of consumers, or in the top 40%, depending on the variance in consumption.

### **The Present Studies**

Attributions regarding health eating can involve both product evaluations and one’s own consumption. We predict that judgments regarding the attributes of products will be made relative to the rank position of the product amongst others that are (a) either present in the actual environment, or (b) retrieved from memory. Similarly, we predict that judgments of one’s own consumption will be made relative to the rank position of one’s consumption amongst other people’s, again involving individuals present in the environment or retrieved from memory. Studies 1a and 1b manipulate the context in which products or people’s consumption is presented to experimentally test our predictions regarding rank based contextual effects. Study 2 tests whether the same process occurs in the absence of experimentally provided contextual cues, where one’s own consumption is judged relative to a sample of other people retrieved from memory. Finally, Study 3 presents a proof-of-concept intervention, testing whether telling people the rank of the consumption relative to others may improve the effectiveness of a social norms intervention relative to the normal presentation of information which focuses on providing nutritional information or how one’s consumption differs from the mean.

#### **Study 1a: Evaluating Food Healthiness**

Study 1a tests (a) whether people's evaluation of the healthiness of a given ready meal (or cereal) will depend on what other products are viewed at the same time (as on a supermarket shelf or in an advertisement), and (b) whether rank- or mean-based models can explain how these contextual effects occur. To achieve this, we experimentally manipulated the rank position of a given product content of salt, fat, and sugar relative to other products viewed at the same time. The rank principle was tested for three different products, in order to increase the generalizability of the results. Participants considered either the salt or fat content of ready meals, or the content of sugar for different cereals. An additional outcome variable measured calorific content; however, as data for the latter showed ceiling effects, they were not further analyzed.

## **Method**

**Participants.** A total of 72 undergraduate students (47 females) from a large public university volunteered to take part in this study. Participants' ages ranged from 18 to 49 years ( $M = 21.65$ ,  $SD = 2.94$ ). Students were enrolled in a variety of undergraduate courses; most students were White (88.33%).

**Design and procedure.** Participants filled in a 3-page questionnaire, one for each different product type (i.e., first ready meal, second ready meal and cereals). On each page, 11 different food items were presented, each with a different content of either salt, fat or sugar (the latter for cereals). For each item, participants were asked to rate its healthiness, on a 1 ("Very healthy") to 7 ("Very unhealthy") Likert scale.

To test DbS and ALT, the distribution of the 11 amounts of content of each substance was manipulated between subjects in order to create two different distributions of the stimuli. Different participants also saw two other distributions of stimuli to examine hypotheses not examined in the present paper; the results from these are not reported here. The substance content in these two distributions is different, with the exception of five amounts which are

presented in both the first (distribution A or unimodal) and second (distribution B or bimodal) distributions (the five ‘common points’).

Consider the example of salt content (in grams) for ready meals (see Figure 1 below). The smallest amount (1.40g; common point 1) was the same in both distributions. This amount has also the same rank position within both distributions (i.e. its rank is equal to 1) and is the same distance from the mean (3.40g) of the set. Therefore no differences in participants’ responses are predicted according to both the absolute and relative accounts of food judgments. The second common point is 2.40g; in distribution A, 2.40g ranks as the 2<sup>nd</sup> lowest (i.e. rank = 2)—while it ranks as the 5<sup>th</sup> in the distribution B (rank = 5). If rank determines people’s evaluations about food, the perceived healthiness of a ready meal containing 2.40g should be lower in distribution B than in distribution A, despite the fact that the content of salt is the same for both ready meals. Also, as 2.40g is the same distance from the distribution’s mean (i.e., it contains 1.00g less than the mean amount of 3.40g), any difference in perceived healthiness for 2.40g cannot be readily explained by reference-level theories.

The distribution mean (3.40g) is the third common point; both its value and its rank position (rank = 6) are the same in both distributions; hence, as for the first common point, no differences in participants’ responses are predicted according to both the absolute and relative accounts of food judgments. Conversely, 4.30g (common point 4) ranks lower in distribution B (rank = 7) than in distribution A (rank = 10)—hence according to rank-based models it should be rated as more unhealthy in distribution A; mean-based accounts on the other hand would predict no differences between the two distributions, as the distance of the fourth common point from the mean of both distributions is the same (i.e. it is 0.90g heavier than the mean). Finally, the largest amount (5.30g) is the fifth common point, and its rank is the same

for both distributions (rank = 11); as for the first and third common points, participants are expected to rate this amount as equally unhealthy in both distributions.

We removed from the analyses participants who responded erratically (see also Melrose, Brown, & Wood, 2013). Specifically, we excluded participants when (1) the Kendall's  $\tau$  coefficient between stimuli (i.e., the 11 food items) and responses (i.e., participants' ratings for the 11 items) was  $< .50$ ; such results were mostly likely due to participants misunderstanding the instructions, for example by (a) assigning progressively low unhealthiness ratings to food items progressively high in content of salt, fat, and sugar – and vice versa – which would cause the  $\tau$  coefficient to be negative and (b) assigning high ratings to high contents and middle ratings for middle contents, only to assign again high ratings for low contents; this would lead  $\tau$  coefficients to be positive but small in size. We also excluded participants if (2) the response range for their ratings within each food domain was  $< 1.00$ ; this criterion ensured that we could remove from the analyses those participants who provided the same unhealthiness ratings regardless of the salt, fat and sugar content under consideration. Application of these criteria resulted in exclusions of 6.94%, 4.17% and 11.11% of participants for salt, sugar and fat contents, respectively. The results were however qualitatively the same when all participants were included in the analyses.

Product scenario (first and second ready meal, cereals) was manipulated within-subjects. The presentation order of the products scenarios was counterbalanced across participants through a Latin square design. The order in which the 11 amounts were presented to each participant was manipulated between-subjects and counterbalanced across products; in the ascending order condition the first of the 11 amounts presented was the smallest, while the opposite was true for the descending order condition.

## **Results and Brief Discussion**

We compared participants' responses for both distributions of content amounts. Figure 2 presents participants' responses for the five amounts that were presented in both conditions (common points). For the 1<sup>st</sup>, 3<sup>rd</sup> and 5<sup>th</sup> common points participants' responses were very similar across the two groups for all the products: this was expected as each amount occupied the same rank position within each distribution type. In line with the rank principle, common point 2 was rated higher (i.e., more unhealthy) in distribution B—where it ranked as 5<sup>th</sup> lowest—than in distribution A, where it ranked 2<sup>nd</sup> lowest. Conversely, common point 4 attracted higher responses in distribution A (rank = 10) than in distribution B (rank = 7). This pattern of results was the same for all three products.

A 5 (within: common point) × 3 (within: substance) × 2 (between: distribution) mixed ANOVA confirmed the observations above. There was a significant main effect of point, as higher amount of salt, fat and sugar were associated with lower healthiness judgments for ready meals and cereals,  $F(4, 244) = 793.64, p < .001, \eta_p^2 = .93$ . More importantly, the interaction between distribution and comparison points was significant,  $F(4, 244) = 15.48, p < .001, \eta_p^2 = .20$ , suggesting that the effects of higher substance content on healthiness ratings depended on each content rank position. This interaction is graphed in Figure 2; as expected, participants' ratings of the 2<sup>nd</sup> and 4<sup>th</sup> common points significantly differed, whereas the ratings of the 1<sup>st</sup>, 3<sup>rd</sup> and 5<sup>th</sup> common points did not. There was also a significant 3-way interaction effect,  $F(8, 488) = 2.70, p = .006, \eta_p^2 = .04$ , meaning that the above interaction effect between common point and distribution type was stronger for salt,  $F(4, 244) = 11.75, p < .001, \eta_p^2 = .16$ , and fat,  $F(4, 244) = 13.49, p < .001, \eta_p^2 = .18$ , than for sugar,  $F(4, 244) = 3.44, p = .009, \eta_p^2 = .05$ .<sup>1</sup>

These results support the hypothesis that, when evaluating the healthiness of food products depending on their content, people make judgments in relative terms in ways predicted by the rank principle (Stewart et al., 2006). Other holistic features of the context

(e.g., content average) did not play a role, thus no empirical support was observed for alternative theories such as reference-level approaches (ALT; Helson, 1947).

### **Study 1b: Perceived Health Risks**

Study 1a showed that the perceived healthiness of a product is heavily determined by the information available in the decision-making context. Study 1b, using the same procedure, extended these findings to another important category of food-related judgments — the perceived health risk as a result of food intake. Participants were asked to rate the perceived likelihood that 11 different people would suffer a health-related illness (e.g., a stroke) as a result of a daily intake of different amounts of salt, fat and sugar. As in Study 1a, different participants also saw two other distributions of stimuli to examine hypotheses not examined in the present paper; the results from these are not reported here.

### **Method**

**Participants.** A total of 42 undergraduate students (26 females) from a large public university in the UK volunteered to take part in this study. Participants' ages ranged from 18 to 41 years ( $M = 20.10$ ,  $SD = 4.53$ ). Students were enrolled in a large variety of undergraduate courses; roughly equal proportions were in either their 1<sup>st</sup>, 2<sup>nd</sup> or 3<sup>rd</sup> year of study. The majority of participants were of White background (64.29%), while smaller proportions were of Chinese (14.29%) and Indian (11.90%) backgrounds.

**Design and procedure.** The method and procedure were similar to those of Study 1a. The participants' task was to rate the likelihood that each of 11 different people—who differed for the daily intake of either salt, fat or sugar—would have a myocardial infarction (for fat), a stroke (for salt), or a generic health difficulty (for sugar). For example, for fat intake, participants indicated the likelihood (as a percentage) in response to the question: “For each person, please indicate the % chance that each person would suffer a myocardial

infarction (commonly known as a heart attack)”. As in Study 1, the comparison between distributions A and B allowed us to test the rank principle (see Table 1 below).

As in Study 1a, participants who provided erratic responses were excluded from the analyses (9.52%, 4.76% and 7.14% for salt, sugar and fat conditions, respectively), as either (a) the Kendall’s  $\tau$  coefficient between stimuli and responses was  $< .50$  or (b) the response range for their ratings within each question scenario was  $< 5$  (i.e. 5% of the possible range of 100). Again, the results were the same when all participants were included in the analyses.

### **Results and Brief Discussion**

Figure 3 presents the estimated percentage likelihood of health-related risks due to food intake for the five common points. As in Study 1a, the 1<sup>st</sup>, 3<sup>rd</sup> and 5<sup>th</sup> common points attracted similar responses; moreover, the same interaction as in Study 1a was observed, whereby the 2<sup>nd</sup> common point attracted higher ratings in distribution B, while the 5<sup>th</sup> common point was rated as riskier in distribution A than in distribution B.

A 5 (within: common point)  $\times$  3 (within: substance)  $\times$  2 (between: distribution) mixed ANOVA confirmed the observations above. As in Study 1 there was a significant main effect of point,  $F(4, 128) = 199.74, p < .001, \eta_p^2 = .86$ . More importantly, the interaction between distribution and comparison points was significant,  $F(4, 128) = 4.82, p = .001, \eta_p^2 = .13$ , confirming that the effects of higher substance content on healthiness ratings depended on each content’s rank position. This interaction is graphed in Figure 3; the 95% confidence intervals for a group that do not bound the mean of the other group indicate statistically significant differences—hence, as expected, participants’ ratings of the 2<sup>nd</sup> and 4<sup>th</sup> common points significantly differed, whereas the ratings of the 1<sup>st</sup>, 3<sup>rd</sup> and 5<sup>th</sup> common points did not. As the 3-way interaction was not significant,  $F(8, 256) = 1.02, p = .418$ , there was no evidence that the effect differed across the three substances.<sup>2</sup>

The results of Study 1b support the conclusions drawn in Study 1a, as the rank principle determined how participants perceived the risk of consuming given amounts of food. No support was observed for either an absolute approach or for reference-level accounts of judgment.

### **Study 2: Social Norms Comparisons**

This study tests whether the rank-based model holds when the context is not represented by other products available at the time of a choice is elicited, but rather the context is the beliefs that the consumer holds in memory. Such a procedure may more closely resemble real-world valuations, in which there may be no comparison items physically present. We elicited participants' perceived distribution of food consumption behavior, specifically how much they think other people consume. As well as looking at a different type of context effects, this study ensures that the previously observed results are not simply an artifact of study design, and that the model holds when questions are based on distributions provided by participants.

In this study, participants reported their weekly consumption of coffee, chocolate, and pizza as well as their attitudes towards consumption. However, as weekly consumption quantities for the latter were very small, data for pizza consumption were not included. We elicited from each participant their beliefs about the social distribution of amounts of consumption of each product. It was hypothesized that individuals' (likely often erroneous) beliefs about their ranked position of their consumption of the products—rather than their beliefs about mean consumption levels, as would be predicted by ALT—will predict their concern about coffee and chocolate consumption.

### **Method**

**Participants.** A total of 201 undergraduate students (138 females) from a large public university volunteered for the study. Participants' ages ranged from 18 to 44 years ( $M =$

20.78,  $SD = 2.43$ ); 77.11% were of White ethnic origin, followed by Chinese (5.97%) and Indian (4.98%) ethnicities.

**Design and procedure.** Participants filled in questionnaires individually. Participants first reported their weight and height, to enable computation of body-mass index (BMI). Next, a task elicited participants' beliefs about the social distribution of consumption of each of two products. There are different ways to elicit probability distributions (Manski, 2004); here, based on pilot work to establish the easiest method for students, we asked participants to estimate different percentiles points of the distribution (Melrose et al., 2013). Nine questions were phrased as follows: "The highest consuming  $x\%$  of students drink more (eat more) than \_\_\_ cups of coffee (bars of chocolate) per week", where  $x$  had values of [10, 20, 30, 40, 50, 60, 70, 80, and 90]. Participants provided estimates for each of the nine percentile points. Results were checked using the same exclusion criteria as in previous studies, and data from approximately 20% of participants were excluded from the analyses (40 for coffee and 41 for chocolate). As before, results were qualitatively the same when all participants were included in the analyses.

Next, participants reported how many cups of coffee (chocolate bars) they drank (ate) per week on average ('own consumption'). Finally, concern about own consumption was measured. Participants answered two questions: "How high do you think your consumption of coffee (chocolate) is" on a 1 ("Very low") to 7 ("Excessive") point scale; and "How concerned are you with your level of coffee drinking (chocolate eating)?" on a on a 1 ("Not at all concerned") to 7 ("Very concerned") point scale.

**Statistical analyses.** To compute the rank position of each participant within what she believed to be the distribution of consumption, we fitted a cumulative distribution function, separately for each participant, to the 9 percentile estimates for each product. We chose either a lognormal function or a linear function depending on which fitted best. We then computed

for each participant (a) the mean of the elicited cumulative distribution ('subjective mean') and (b) the relative rank position of each participant's consumption within such distribution ('subjective rank').

### Results and Brief Discussion

**Estimating coffee and chocolate consumption.** Students greatly overestimated the consumption of others (see Figure 4 below) and underestimated their own rank position within the true consumption distribution. On average, students believed that 70% of students consumed more of the two products than they did themselves: Subjective rank was well below .50 for coffee ( $M = .35$ ,  $SD = .25$ ; interquartile range,  $IQR = [.15, .52]$ ) and chocolate ( $M = .25$ ,  $SD = .20$ ;  $IQR = [.10, .33]$ ).

The large variation in beliefs about other people's consumption is exemplified in Figure 6, which shows the beliefs about the number of bars of chocolate consumed by other students for participants 68 and 190. Although participant 190 consumed more bars of chocolate per week than did participant 68, she believed that only 20% of students consumed less than she did herself (whereas the subjective rank for participant 68 was above .50), and—as predicted—she reported lower concern about her own chocolate consumption.

**Predicting attitudes towards food consumption.** We ran ordinal regression analyses to predict students' attitudes towards their diet (i.e., how high they considered their consumption, and how concerned they were about it). Predictors were 'subjective rank', 'subjective mean', and 'own consumption'; control variables were BMI, gender (1=Females, 2=Males), and age.

Table 2 shows that the results were as predicted by the rank-based models. Judgments about own consumption were predicted by own consumption level and subjective rank, but not by beliefs about mean consumption within the social comparison group (with the exception of concern about own coffee consumption). An additional effect was observed for

gender, whereby male students generally reported lower level of concern for chocolate and coffee consumption than did female students.<sup>3</sup> The regression models were further analyzed as suggested in Andraszewicz et al. (2014). In particular, the full regression model (*f*) including the critical predictor (subjective rank) was compared to the constraint model (*c*) that excluded it (i.e., the model that included only own consumption, subjective mean, age, gender and BMI as the predictors) for each of the four outcome variable. The analyses showed that for chocolate, the evidence in favor of the full model was anecdotal ( $BF_{fc} = 0.57$  and  $BF_{fc} = 0.34$  for high consumption and concern about consumption, respectively). On the other hand, the results for coffee consumption provided moderate evidence for the full model ( $BF_{fc} = 0.14$  and  $BF_{fc} = 0.12$  for high consumption and concern about consumption, respectively). Thus for chocolate consumption the evidence in favor of rank-based models was somewhat weaker than for coffee consumption, although the qualitative pattern of data was similar. It may be that participants' prior knowledge about daily recommendations for sugar intake may have moderate the observed contextual effect, although further research could address this issue directly.

We also analyzed the moderating roles of gender and BMI on the relative rank effects we observed; however, no moderating effects were found. For instance, we analyzed whether gender moderated the effects of rank on participants' perception of the healthiness of their own consumption of chocolate and coffee. We did so by adding to the regression equation the term gender by subjective rank, this interaction term never reaching significance (all  $ps > .185$ ). Similarly, when the term BMI by subjective rank was entered in the equations, no moderation effects were observed (all  $ps > .116$ ). At the same time, when the interaction terms were entered into the equations, no significant changes were observed for the coefficients of three main predictors (i.e., subjective rank, subjective mean and own consumption). These findings indicate that the same pattern of findings is observed regardless

of gender and BMI. The results suggest that judgments about food consumption, and attitudes towards diet were determined by how participants thought their consumption ranked within what they believed other students consume, with an additional contribution of own consumption level.

### **Study 3: Behavioral Nudging**

Study 2 showed that concern about food intake is driven by individuals' beliefs about where their own consumption ranks amongst others. Study 3 tests whether providing information about the rank of participants' own consumption levels influences their food preferences more strongly than information about mean consumption. We used an incentive compatible paradigm (the Becker-DeGroot-Marschak procedure, BDM; Becker, DeGroot, & Marschak, 1964) to examine participants' willingness to pay (WTP) for relatively healthy and unhealthy products. Thus, the present study builds on the previous ones in order to ascertain whether attitudes and hypothetical choices and judgments can predict actual behavior.

The present study also tests whether rank-based models (Parducci, 1965; Stewart et al., 2006) can predict actual behavioral outcomes, which here have been operationalized as monetary choices for food products. We also test whether manipulating the way in which social normative information (i.e., the actual level of consumption of other students) is provided has an effect in people's actual choices for food products. In line with DbS and rank-based approaches more generally, we hypothesize that providing rank-based normative information (e.g., "90% of people consume less chocolate than you do") will be more effective than mean-based 'nudging' ("you eat 5 bars of chocolate per week; on average, other people consume 3 bars per week") in influencing consumers' choice of food products.

### **Method**

**Participants.** Ninety five participants (55 females) who reported any consumption of either coffee or chocolate were included. Participants were recruited from a large public

university campus; ages ranged from 18 to 49 years ( $M = 22.18$ ,  $SD = 4.48$ ). Participants received £5 (at the time of the study, £GBP 1 = \$USD 1.58) in exchange for their participation, minus any cost incurred in the BDM procedure.

**Design and procedure.** Participants first (a) performed the distribution elicitation task for either coffee or chocolate consumption and (b) indicated their own weekly consumption of either product. These tasks were counterbalanced between participants. Distribution elicitation questions were worded as follows: “The highest consuming  $x\%$  of students drink (eat) \_\_\_ cups of coffee (bars of chocolate) per week”, where  $x$  values were [10, 20, 30, 40, 50, 60, 70, 80, and 90]. Using the same criteria as in the previous studies, prior to analysis we removed data from 12 participants who responded erratically (12.63% of the total), although including all participants in the analyses did not affect the results.

Participants were then provided with information relating to their coffee (or chocolate) consumption. There were three conditions. In a “mean product information nudging” condition participants were told the average calorific (chocolate) or caffeine (coffee) content of the foods. In a “mean consumption nudging” condition participants were told what they believed the average consumption to be (based on the estimates they provided in the elicitation task), and what the actual average consumption was (based on normative data previously collected from 263 students at the university under consideration). In the final condition, “rank-based nudging”, participants were told where they believed they ranked among the university student population for coffee (or chocolate) consumption, and what their actual rank position was; the former was the participant’s percentile position within their inferred distribution of consumption, while the latter was the percentile position within the actual distribution of consumption (estimated from the normative study).

Participants were then asked how much they would pay to purchase each of two products. Depending on group, participants were told that the pairs of products on sale were

either (a) a standard 49g (1.73 oz.) Cadbury® bar of chocolate and a red apple, or (b) a 250ml (8.45 fl. oz.) bottle of Starbucks® Mocha Frappuccino and a 250ml Tropicana® orange juice. The store prices of the products were reasonably matched: The bar of chocolate cost 59p (while the apple cost 50p) and the frappuccino cost £1.49 (while the orange juice cost £1.29).

Participants first stated how much (in pence) they would be willing to pay for each of the products. A random number generator then determined whether it was the chocolate (frappuccino) or the apple (fruit juice) that was on sale. A random number generator then determined the price. If the random price was lower or equal to their bid, participants had to buy the (randomly selected) product for the randomly generated price. Otherwise, if the random price was higher, no transaction took place. It was clearly explained to participants that this procedure ensured that it was best for them to truthfully reveal the price they were willing to pay for each product.

## **Results and Brief Discussion**

**Estimating coffee and chocolate consumption.** As in Study 2, participants in the rank-based condition underestimated their own rank position; on average, the difference between their subjective rank position (their rank position within the inferred distribution) and their actual rank position (their rank position within the normative distribution) was around 30% ( $M = -.33$ ,  $SD = .18$ ,  $IQR = [-.26, -.44]$ ). Similarly, participants in the mean consumption feedback condition overestimated other people's mean weekly consumption by around three cups of coffees (or bars of chocolate;  $M = 3.02$ ,  $SD = 4.86$ ,  $IQR = [0.17, 4.87]$ ; see Figure 5 below). As in Study 2, participants underestimated their rank position within the normative distribution of consumption. The empirical data do not allowed us to us to determine whether people incorrectly reported their own consumption (underestimating it) or they misestimated other people's behavior (overestimating it), although evidence in the domain of alcohol consumption suggests the latter (Perkins, Haines, & Rice, 2005). However,

our result provides support for the need for social norms interventions which tell people about their actual rank position within the social distribution of food consumption.

**The effects of feedback information on bidding behavior.** We computed the “healthy bid ratio” (i.e., the bid for the healthy product over total bid) to minimize the effects of inter-individual differences in bidding behavior. A bid ratio below .50 shows a preference for the unhealthy food. The bid ratio was considerably higher for participants in the rank group ( $M = .53$ ,  $SD = .17$ ) than for participants in the mean product information ( $M = .40$ ,  $SD = .14$ ) and mean consumption conditions ( $M = .39$ ,  $SD = .13$ ). A one-way between subjects ANOVA confirmed these observations,  $F(2, 82) = 7.73$ ,  $p < .001$ ,  $\eta_p^2 = .16$ ; follow-up  $t$ -tests (with Bonferroni adjustments) revealed that the ratio bid was significantly higher for the rank group compared to mean product and mean consumption information groups (both  $ps < .01$ ;  $d = 0.81$  and  $d = 0.93$ , respectively) – the latter two did not differ ( $p > .95$ ).<sup>4</sup>

These results show that participants in the rank group were willing to spend relatively more for a healthier product compared to the other two groups. Moreover, providing mean consumption feedback information had no effect on WTP.

Finally, to test the effects of feedback information on bidding behavior, we examined whether, in the rank-based condition, greater underestimation of an individual’s own rank position was associated with greater willingness to pay for the healthy option. Regression analysis was used to predict the healthy bid ratio from both the feedback difference (i.e., the difference between the actual values of average or rank consumption and participants’ estimates of average consumption or their own rank position) and participants’ own consumption—separately for the rank- and mean- consumption based groups. Consistent with the rank-based model of judgment, feedback difference significantly predicted bid ratio in the rank-based feedback condition,  $\beta = .40$ ,  $p = .036$ , but not in the mean consumption feedback condition,  $\beta = -.12$ ,  $p = .609$ —the two coefficients differing significantly,  $z = 1.77$ ,  $p = .039$

(1-tailed). Thus, in the latter feedback condition, participants' bidding was independent of the feedback they received; conversely, in the rank group, the higher the underestimation of own rank position, the higher the willingness to pay for the healthy option (compared to the unhealthier one).<sup>5</sup> The procedure we used aimed to elicit the true prices participants were willing to pay for each product; thus, the healthy bid ratio reflected participants' preferences at the time of testing. Nonetheless, it remains an empirical question whether the same results would be observed in a forced choice task paradigm, and we hope that future research will aim to replicate these findings using different decision-making paradigms.

The results of this study show that feeding back information about rank consumption is an effective way to reduce people's willingness to pay for unhealthy products. Providing information about average content of calories (or caffeine), or about average consumption, did not exert any effect on bidding behavior.

### **General Discussion**

The environmental context in which a food is encountered influences how that food is evaluated and whether it is purchased (Chandon & Wansink, 2012). This occurs because foods are not naturally chosen for purchase or consumption in isolation, but are compared (automatically) to similar foods present in the immediate proximity or retrieved from memory. Thus, different immediate or remembered contexts would be expected to lead to different food evaluations and purchase. Previous studies of context effects in food related judgments have however not provided a model of the precise cognitive mechanisms involved, and have rarely tested how results could inform interventions to improve health food behavior. Thus, we tested the ability of rank-based models of judgment and decision-making (DbS and RFT) to explain the cognitive processes through which context impacts on food evaluation and choice—against the ability of a reference-level approach (ALT) to account for the same effects. Studies 1a and 1b showed that participants' perception of the harmfulness of

food products and risk as a consequence of food intake is highly rank-dependent. Study 2 showed that participants' concern about their own consumption of relatively unhealthy products depends both on (a) the consumption level and (b) how they believe their consumption ranks amongst others'. Study 3 showed that a rank-based nudge affected purchasing behavior. Although participants held inaccurate beliefs about both (a) the average consumption of relatively unhealthy food and (b) their own rank position among other consumers, only addressing the latter had a positive effect on participants' willingness to pay for a healthy food option.

Through understanding the precise cognitive mechanisms through which behavioral decisions are made, interventions can be precisely targeted to work with the grain of human nature, and products can be designed and packaged to have the biggest impact on healthy consumer behavior. The results of Study 3 may inform social norm-based feedback components of multi-component and theoretically driven social marketing intervention studies: We showed that the provision of rank-based social norm information may be a more effective nudge than provision of mean-based information (Schultz et al., 2007). These results suggest ways in which psychological principles can be used (e.g., via "nudging") to increase consumption of healthy foods. For example, personalized rank-relevant information could be provided to the wider public through an internet-based campaign. Respondents would be invited to fill in a short on-line questionnaire; the items would mimic those used for Study 3 and would refer to the respondents' consumption of particularly unhealthy foods. Normative data would have been collected on these particular foods and it would be used to easily extrapolate rank-relevant information for each respondent as s/he enters his/her answer about food consumption. We believe that this way of delivering personalized, rank-relevant information could prove a rather effective way in "nudging" people into healthier decision-

making about food consumption—at least to a greater extent than general and impersonal information campaigns.

The applicability of the present rank-based model to the domain of judgments about food may however be limited by several factors. First, in Study 2 participants provided what they believed to be the distributions of others' consumption before they answered questions about their own consumption. It is possible that this primed people to make judgments in a more relative way than normal. However, this task order was preferred as it was thought to parallel how judgments are made in practice; consumption occurs within a context and then judgments are made, rather than judgments creating contexts (at least on the occasion that judgments are made). Conversely, if people made the judgments first, they may have been led to estimate the distributions in a way that is consistent with the judgments they made, which was considered to be a greater potential problem. However, as the tasks were counterbalanced in Study 3, we could analyze the influence of order effects and we found none. We note nonetheless the limitation and that future work should address this issue further.

In addition, in some of the reported studies it was necessary for quantities to be present in the decision-making context as participants had to provide evaluations based on numerical values (e.g., the likelihood of health risks due to food intake or the amount of cups of coffee drunk each week). In many real world contexts – such as when browsing supermarket shelves – people are indeed exposed to quantities and different options in a similar fashion and they have to make decisions accordingly. We also note that the same relative rank effects emerged in experimental contexts where people had to make judgments based on information available in the immediate decision context and on information retrieved from memory.

Conclusions

Our key theoretical advance is the application of rank-based models of contextual and social comparisons to consumer research. We demonstrated the utility of this approach through direct comparison with the mean-based reference level model which is normally applied in consumer research to explain how people make relative comparisons. In Studies 1a and 1b we held constant across conditions the distance between individual people/products and the mean of those in the comparison set, such that any observed differences between groups would not be predicted by the reference level account. In the cross-sectional Study 2 we controlled for distance for the mean statistically, and in the third, intervention, study we directly tested the differences in purchase behavior resulting from telling a person their rank position amongst others rather than how they compare to the average person.

Social norms interventions – information on how one’s undesirable behavior compares to other – are very widely used across social marketing and health research more generally. However, their effectiveness is limited. Our key applied advance is the suggestion that social norms interventions can be improved through a minor change of presentation of information, focusing on providing rank position of one’s behavior rather than how one’s behavior compares to others. This subtle reframing provides information in ways in which people naturally process it and as such increases effectiveness substantially.

Finally we present a theoretical advance for the DbS model itself. Previous research on DbS has often used subjective evaluations rather than actual behavior as the outcome. Often this is appropriate as such evaluations are the variable of interest. However, criticisms have previously been made that the observed effects are simply artifactual, in that they represent differences in how the rating scales are used (e.g., “label stretching”) due to different contexts, rather than actual differences in true evaluation. In Study 3 we address this concern through showing that the model predicts actual behavior. Study 3 also illustrates the potential real-world application of rank-based models: It is argued that social norms

intervention can maximize their effectiveness by providing rank-relevant information (e.g., “you are in the most unhealthy 10% of eaters”) rather than information relative to average behaviors (e.g., “you consume 500 calories more than the average person”).

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## Footnotes

<sup>1</sup> When all participants were included in the analyses, the following significant effects were observed: (a) a main effect of point,  $F(4, 280) = 516.16, p < .001, \eta_p^2 = .88$ , (b) an interaction between point and distribution,  $F(4, 280) = 7.73, p < .001, \eta_p^2 = .10$  and (c) a 3-way interaction,  $F(8, 560) = 2.21, p = .025, \eta_p^2 = .03$ . However, the interaction between point and distribution was significant for salt,  $F(4, 280) = 8.14, p < .001, \eta_p^2 = .10$ , and fat,  $F(4, 280) = 10.15, p < .001, \eta_p^2 = .13$ , but not for sugar,  $F(4, 280) = 0.56, p = .692$ .

<sup>2</sup> The same results were observed when all participants were included in the analyses, as we observed (a) a main effect of point,  $F(4, 160) = 163.88, p < .001, \eta_p^2 = .80$ , (b) an interaction between point and distribution,  $F(4, 160) = 3.62, p = .007, \eta_p^2 = .08$ , while (c) the 3-way interaction was not significant,  $F(8, 320) = 0.33, p = .953$ .

<sup>3</sup> Results were similar when all participants were included in the analyses. For coffee, for the dependent variables (a) concern about consumption, subjective rank,  $B = 5.21, Wald = 26.94, p < .001$ , was the only significant predictor (all other  $ps > .152$ ) and (b) high consumption, subjective rank,  $B = 5.02, Wald = 24.34, p < .001$ , own consumption,  $B = 0.31, Wald = 38.88, p < .001$ , and subjective mean,  $B = 0.13, Wald = 4.44, p = .035$ , were all significant predictors, while no other variable was (all  $ps > .475$ ). For chocolate, for the dependent variables (a) concern about consumption, subjective rank,  $B = 1.79, Wald = 6.84, p = .009$ , and own consumption were significant predictors,  $B = 0.03, Wald = 7.45, p = .006$ , while subjective mean was not,  $B = 0.01, Wald = 1.02, p = .312$ ; BMI,  $B = -0.03, Wald = 4.38, p = .036$ , and gender,  $B = -0.80, Wald = 7.33, p = .007$ , were significant predictors, too; and (b) high consumption, both subjective rank,  $B = 5.11, Wald = 44.40, p < .001$ , and own consumption,  $B = 0.07, Wald = 17.37, p < .001$ , were significant predictors, while the subjective mean was not,  $B = 0.01, Wald = 2.47, p = .116$ ; gender was the only other significant predictor,  $B = -0.79, Wald = 7.34, p = .007$ .

<sup>4</sup>As for the previous studies we performed additional analyses, including all participants; the main effect of feedback type on bid ratio was significant,  $F(2, 92) = 7.33, p = .001$ ;  $t$ -tests with Bonferroni adjustments confirmed that the bid ratio was higher for the rank group ( $M = .51, SD = .17$ ) compared to mean-based feedback ( $M = .39, SD = .12; p = .007$ ) and control group ( $M = .40, SD = .12; p = .003$ )—the latter two groups did not differ ( $p > .97$ ).

<sup>5</sup>We ran further analyses in order to rule out the possibility that task order effects might have moderated the above findings. The results showed order did not influence participants' misestimations, as no main effect of the variable order was detected on misestimations, nor its interaction with the feedback type (both  $F_s < 1$ ). The same was true when the same analysis was run on the healthy bid ratio: Neither the main effect of order,  $F(1, 77) = 1.27, p = .263$ , nor the interaction order by feedback type ( $F < 1$ ) was significant. Lastly, when the variable order was entered in the last regression analyses on healthy bid ratio with own consumption and feedback difference as the predictors, it did not significantly predict nor interact with any of the predictors (all  $p_s > .07$ ).

Table 1

*Average daily consumption of salt, fat and sugar (in g) for distribution A (unimodal) and distribution B (bimodal). Underlined amounts represent the five common points between the distributions.*

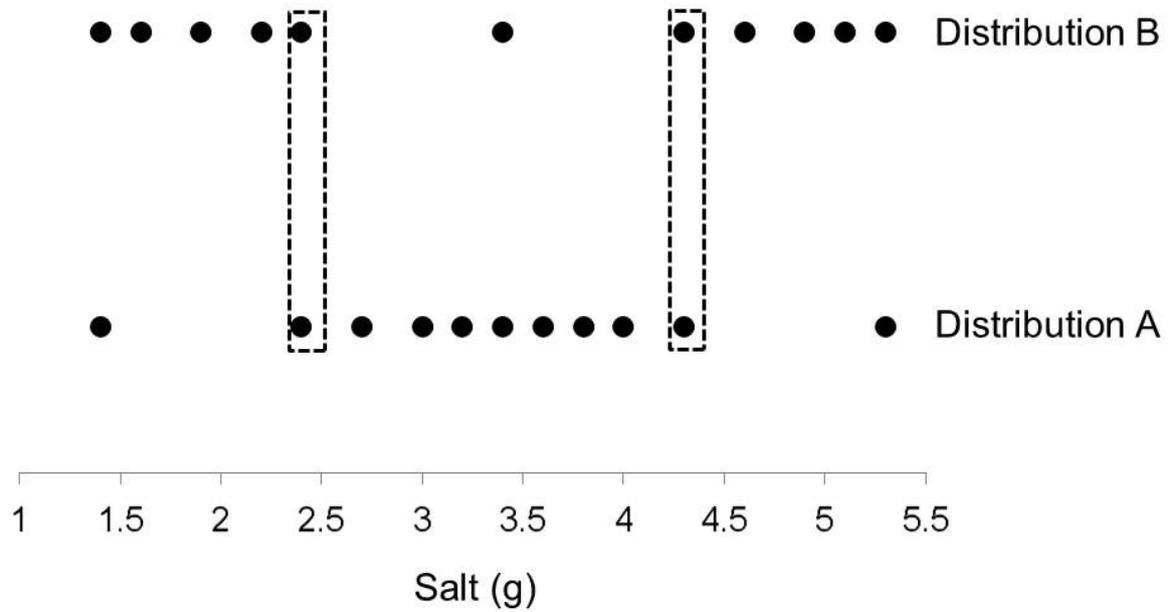
Substance intake					
Salt		Fat		Sugar	
Distribution					
A	B	A	B	A	B
<u>0.40</u>	<u>0.40</u>	<u>10.00</u>	<u>10.00</u>	<u>10.00</u>	<u>10.00</u>
<u>2.80</u>	0.90	<u>50.00</u>	18.00	<u>66.00</u>	22.00
3.50	1.40	58.00	26.00	82.00	34.00
4.00	2.10	66.00	38.00	94.00	50.00
4.50	<u>2.80</u>	74.00	<u>50.00</u>	106.00	<u>66.00</u>
<u>5.00</u>	<u>5.00</u>	<u>82.00</u>	<u>82.00</u>	<u>118.00</u>	<u>118.00</u>
5.50	<u>7.20</u>	90.00	<u>114.00</u>	130.00	<u>170.00</u>
6.00	7.90	98.00	126.00	142.00	186.00
6.50	8.60	106.00	138.00	154.00	102.00
<u>7.20</u>	9.10	<u>114.00</u>	146.00	<u>170.00</u>	214.00
<u>9.60</u>	<u>9.60</u>	<u>154.00</u>	<u>154.00</u>	<u>226.00</u>	<u>226.00</u>

Table 2

*Regression coefficients for the analyses on attitudes towards own consumption of chocolate and coffee.*

	Estimate	SE	Wald	<i>p</i>		Estimate	SE	Wald	<i>p</i>
<i>Coffee</i>					<i>Chocolate</i>				
<i>Consumption, High</i>					<i>Consumption, High</i>				
Subjective Rank	6.49	1.50	18.77	<.001***	Subjective Rank	1.86	1.23	2.29	.130
Own consumption	0.12	0.06	4.92	.027*	Own consumption	0.70	0.11	37.37	<.001***
Subjective Mean	0.01	0.04	0.10	.754	Subjective Mean	<0.01	0.02	0.19	.662
BMI	0.02	0.02	0.52	.471	BMI	<0.01	0.02	<0.01	.982
Gender	-0.27	0.42	0.40	.526	Gender	-0.92	0.36	6.55	.011*
Age	0.01	0.10	0.02	.888	Age	-0.04	0.06	0.35	.556
<i>Consumption, Concern</i>					<i>Consumption, Concern</i>				
Subjective Rank	4.36	1.31	11.08	<.001***	Subjective Rank	2.79	1.18	5.58	.018*
Own consumption	0.06	0.04	2.78	.096•	Own consumption	0.17	0.08	4.41	.036*
Subjective Mean	0.07	0.03	4.55	.033*	Subjective Mean	<0.01	0.02	0.16	.689
BMI	-0.01	0.03	0.04	.840	BMI	-0.03	0.02	1.73	.189
Gender	-0.91	0.47	3.79	.052•	Gender	-1.48	0.37	15.91	<.001***
Age	0.07	0.10	0.44	.509	Age	-0.10	0.07	1.88	.171

*Note.* BMI = Body-Mass Index; \*\*\* significant at 0.1% level; \*\* significant at 1% level; \* significant at 5% level; •  $p < .10$ .



*Figure 1.* Illustration of distribution A (bottom line of black dots) and distribution B (upper line of black dots) constructed to test relative rank effects. Highlighted amounts are common points 2 and 4, which differ only in their rank position within the distribution and not in their magnitude or in how they differ from the mean of the set.

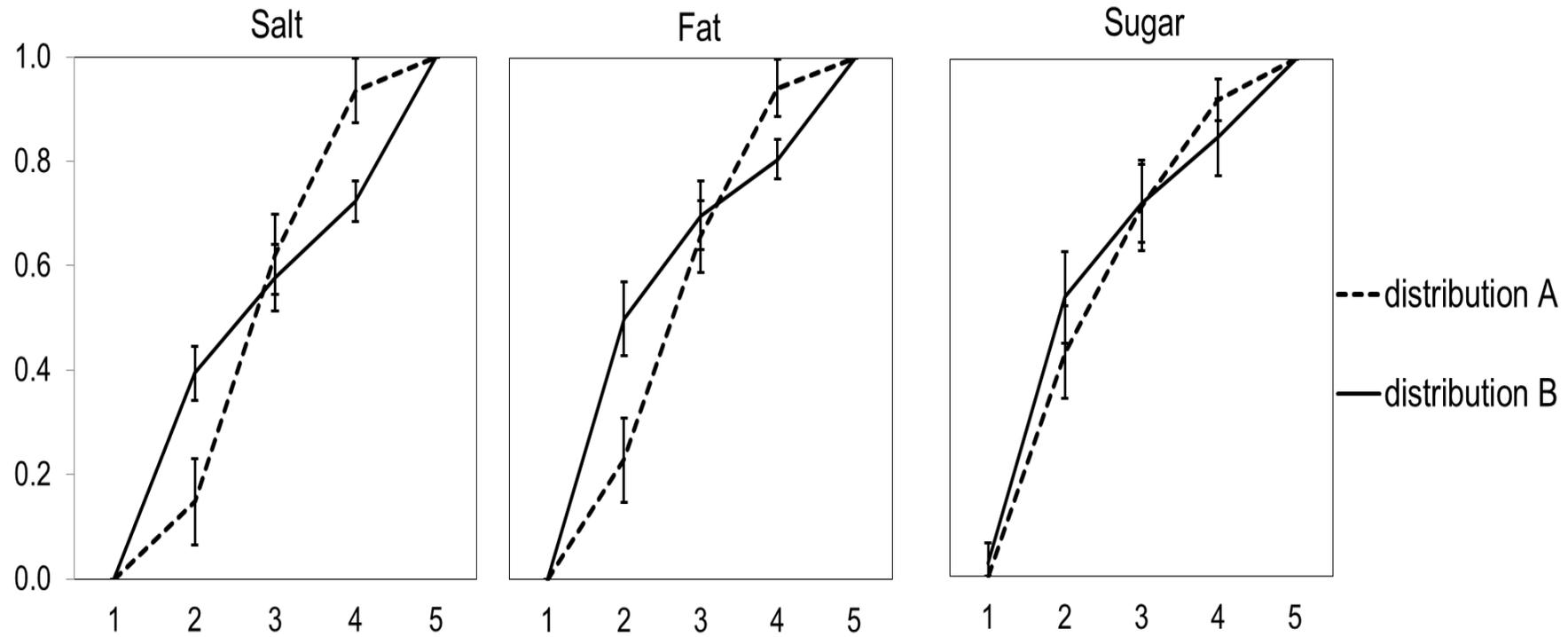


Figure 2. Interactions between distribution type and common points (1 to 5) for the three substances. To take account of individual differences in scale use, all participants' responses were rescaled between 0 and 1 (see Brown et al., 2008). Error bars represent 95% confidence intervals.

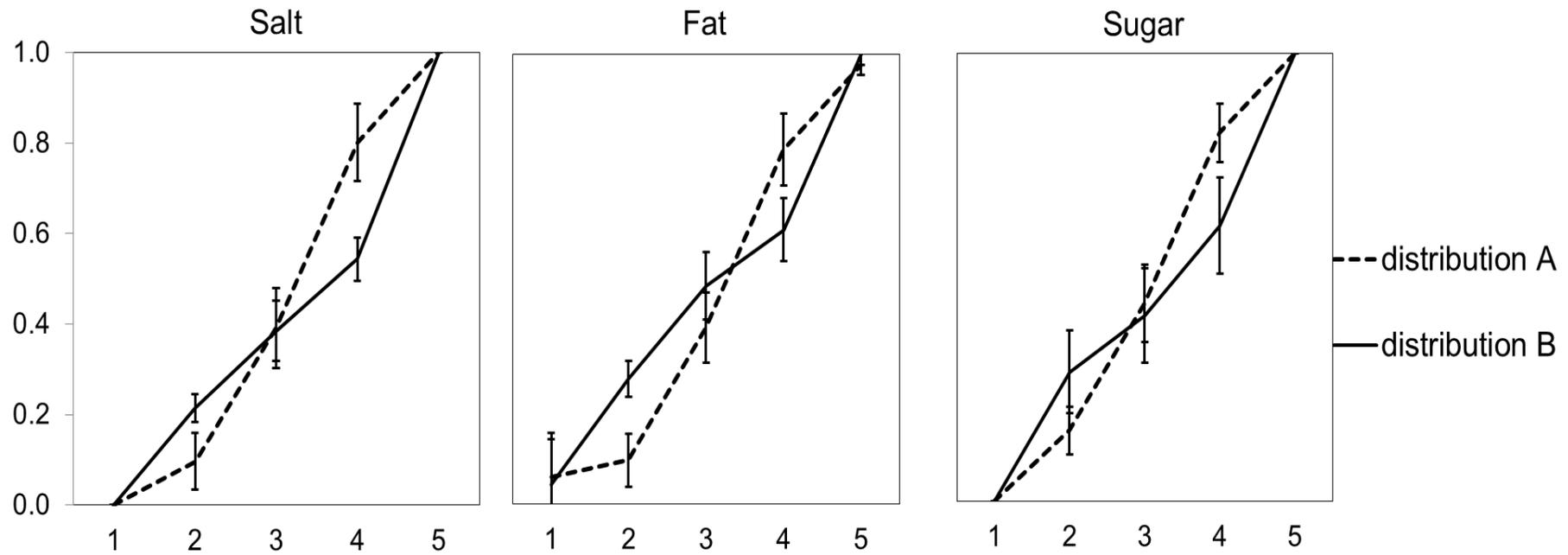


Figure 3. Interactions between distribution type and common points (1 to 5) for the three substances. To take account of the individual differences in scale use, all participants' responses were rescaled between 0 and 1 as in Study 1b. Error bars represent 95% confidence intervals.

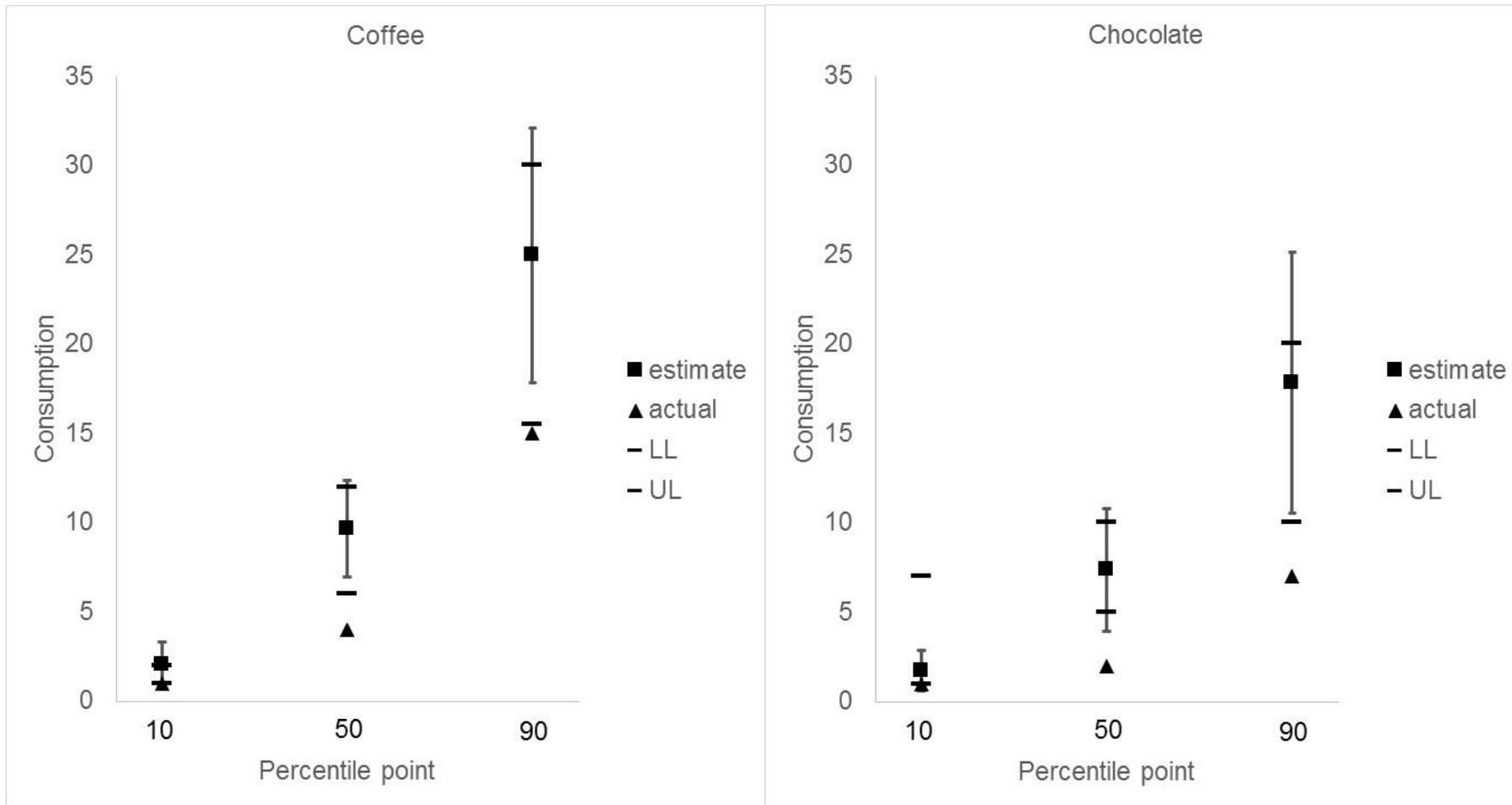


Figure 4. Consumption mean estimates (and SD) and actual values for coffee (cups; on the left) and chocolate (bars; on the right) provided in the probability elicitation task for different percentile points. LL and UL represent the lower and upper limit of the Interquartile Range (IQR). Whiskers present the SD of participants' estimates.

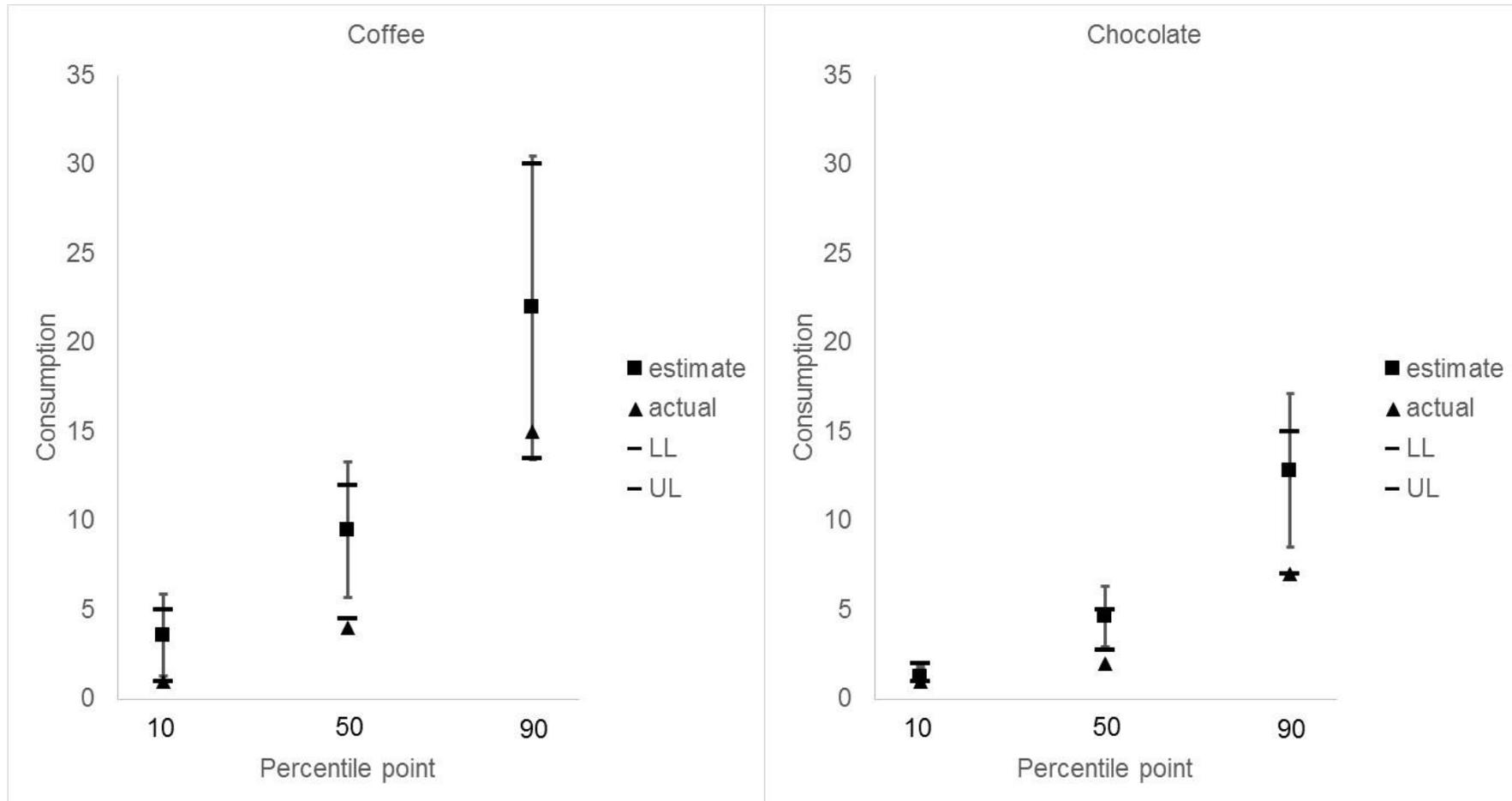


Figure 5. Consumption mean estimates (and SD) and actual values for coffee (cups; on the left) and chocolate (bars; on the right) provided in the probability elicitation task for different percentile points. LL and UL represent the lower and upper limit of the Interquartile Range (IQR). Whiskers present the SD of participants' estimates.

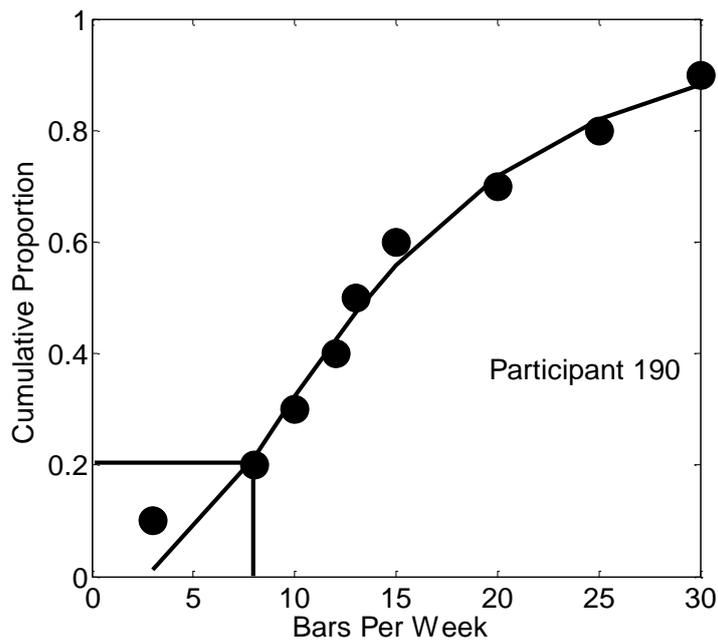
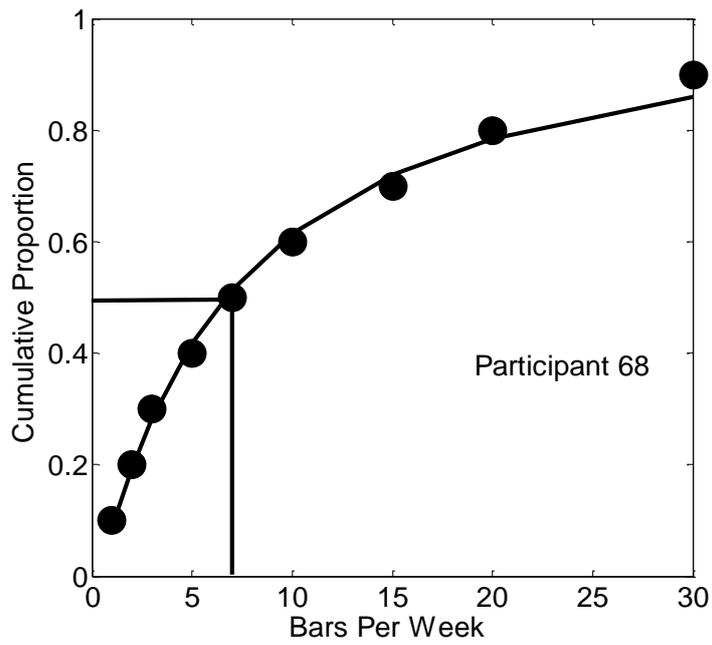


Figure 6. The cumulative distribution (filled circles) along with a best-fit cumulative density function (solid lines) elicited from participants 68 (top panel) and 190 (bottom panel). Vertical lines indicate own consumption, while the horizontal line represents the subjective rank position.